



## On the existence of moments in Cauchy-like distributions induced from the tan function

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### Abstract

In this paper we consider cases of the existence of the moments of functions of random variables supported on a bounded interval. Our attention is restricted to the tan function, as a generalization of the Cauchy distribution which is infact the result of applying this function to a uniformly distributed variable.

**Keywords:** Cauchy distributions, tan function, moments

### Preliminaries

The standard Cauchy distribution can be derived as follows - having a uniformly distributed random variable  $X$  on  $\left(-\frac{\pi}{2}; -\frac{\pi}{2}\right)$  we define  $Y = \tan(X)$ . This distribution is famous for its lack of moment existence. [1] [4] [5]

The following question arises naturally - if we allow any distribution for  $X$ , perhaps with a well-defined probability density or probability mass function, under what conditions at least some moments of the variable  $Y = \tan(X)$  would exist and up to what order? Such questions are important in practice, since real-life modelled variables are generally bounded and all of their moments exist - however, it seems this is not always true for functions of such [3].

As a start, let us consider  $Y = g(X)$  where  $g : I \rightarrow \mathfrak{R}$  is a well-defined strictly increasing diffeomorphism between an interval  $I$  with endpoints  $a$  and  $b$  and  $\mathfrak{R}$ . In terms of the cumulative distribution functions of  $Y$  and  $X$  we have

$$F_Y(x) = P[Y \leq x] = P[X \leq g^{-1}(x)] = F_X(g^{-1}(x))$$

via which we have the form

$$E[Y^n] = \int_{-\infty}^{\infty} x^n dF_Y(x) = \int_{-\infty}^{\infty} x^n dF_X(g^{-1}(x)) = \int_{-\infty}^{\infty} x^n f_X(g^{-1}(x))(g^{-1}(x))' dx$$

Another expression, more suitable for our purposes, is the following

$$E[Y^n] = \int_{-\infty}^{\infty} x^n dF_X(g^{-1}(x)) = \int_a^b g^n(y) dF_X(y) = \int_a^b g^n(y) f_X(y) dy$$

The discrete version of the latter is as follows

$$E[Y^n] = \sum_{y \in X} g^n(y) f_X(y)$$

where  $f$  is the probability mass function of  $X$ . Throughout the rest of the article, we treat  $I = \left(-\frac{\pi}{2}; \frac{\pi}{2}\right)$  and  $g(x) = \tan x$ .

## General analysis for tan and some examples

Firstly, let us consider the discrete variable case. Here it will be useful that absolute convergence behaves quite well with sums.

**Proposition 1.** If the set  $\{\tan y | y \in X\}$  is bounded, then  $E[Y^n]$  exists for all  $n$ .

*Proof.* It suffices to consider  $K_0 = \sum_{y \in X} |\tan^n(y) f_X(y)|$ . If  $M$  is a bound for the given set, then as  $\sum_{y \in X} f_X(y) = 1$  and  $f_X(y) \geq 0$ , we have  $K_0 \leq M^n$ , thus the sum  $E[Y^n]$  is (absolutely) convergent.

Of course, in general this need not be the case.

**Example 2.** Let  $X$  attain the values  $y_i = \arctan(2^i)$ ,  $i \geq 1$ , with probabilities

$f_X(y_i) = \frac{1}{2^i}$  (these are positive and sum exactly to 1). Then the  $k$ -th partial sum of the  $n$ -th moment is  $\sum_{i=1}^k 2^{i(n-1)} \geq \sum_{i=1}^k 1 = k$  and thus  $\tan(X)$  has no moments.

Now we move on to the continuous variable case.

In  $E[Y^n] = \int_{-\frac{\pi}{2}}^{\frac{\pi}{2}} \tan^n(y) f_X(y) dy = \int_{-\frac{\pi}{2}}^{\frac{\pi}{2}} \frac{f_X(y)}{\cos^n y} (\sin^n y) dy$  the  $\sin$  function is irrelevant at the only possibly problematic points  $\pm \frac{\pi}{2}$ . Moreover, by Taylor's theorem we have  $\cos y = \left(\frac{\pi}{2} \pm y\right) + O\left(\left(\frac{\pi}{2} \pm y\right)^3\right)$ . As a consequence we have established the following

**Proposition 3.** The  $n$ -th moment exists if and only if the limits

$$\lim_{R \rightarrow \frac{\pi}{2}} \int_R^{\frac{\pi}{2}} \frac{f_X(y)}{\left(\frac{\pi}{2} - y\right)^n} dy \quad \text{and} \quad \lim_{R \rightarrow -\frac{\pi}{2}} \int_{-\frac{\pi}{2}}^R \frac{f_X(y)}{\left(\frac{\pi}{2} + y\right)^n} dy$$

both exist.

In particular:

**Example 4.** In the standard Cauchy distribution we have  $f_X \equiv \frac{1}{\pi}$ , so neither of the limits above exists and the moments are undefined, as expected.

**Example 5.** If there exist two points  $x_1$  and  $x_2$  with  $-\frac{\pi}{2} < x_1 < x_2 < \frac{\pi}{2}$  and  $F_X(x_1) = 0$ ,  $F_X(x_2) = 1$ , then all moments exist. Indeed, in this case the density is 0 outside of  $[x_1; x_2]$  and thus the above limits are equal to zero.

**Example 6.** If  $f_X(x) = C \left(x + \frac{\pi}{2}\right)^p \left(\frac{\pi}{2} - x\right)^q$  for  $p, q > 0$  and an appropriate normalizing constant  $C > 0$ , then the existing moments are precisely those up to order  $\min(p; q)$ .

**Example 7.** If  $f_X(x) = C \exp\left(-\frac{M}{\left(x + \frac{\pi}{2}\right)^p \left(\frac{\pi}{2} - x\right)^q}\right)$  for  $M, p, q > 0$  and an appropriate normalizing constant  $C > 0$ , then all moments exist.

## References

- [1] Stoyanov, J., Counterexamples in Probability, Third Edition, Dover Publications, Inc., (2013).

- [2] Laha, R. G., An example of a nonnormal distribution where the quotient follows the Cauchy law. Proc. Nat. Acad. Sci. USA 44, (1958) 222-223
- [3] Pitman, E. J. G. and Williams, E. G., Cauchy-distributed functions of Cauchy variates. AMS 38, (1967) 916918
- [4] N. L. Johnson; S. Kotz; N. Balakrishnan, Continuous Univariate Distributions, Volume 1. New York: Wiley., (1994), Chapter 16.
- [5] Feller, William, An Introduction to Probability Theory and Its Applications, Volume II (2 ed.). New York: John Wiley & Sons Inc., (1971), p. 704